**Analysis of Crash Proportion by Vehicle Type at Traffic Analysis Zone Level:**

**A Mixed Fractional Split Multinomial Logit Modeling Approach with Spatial Effects**

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May 2017

Revised for Possible Publication in Accident Analysis and Prevention

**ABSTRACT**

In traffic safety literature, crash frequency variables are analyzed using univariate count models or multivariate count models. In this study, we propose an alternative approach to modeling multiple crash frequency dependent variables. Instead of modeling the frequency of crashes we propose to analyze the proportion of crashes by vehicle type. A flexible mixed multinomial logit fractional split model is employed for analyzing the proportions of crashes by vehicle type at the macro-level. In this model, the proportion allocated to an alternative is probabilistically determined based on the alternative propensity as well as the propensity of all other alternatives. Thus, exogenous variables directly affect all alternatives. The approach is well suited to accommodate for large number of alternatives without a sizable increase in computational burden. The model was estimated using crash data at Traffic Analysis Zone (TAZ) level from Florida. The modeling results clearly illustrate the applicability of the proposed framework for crash proportion analysis. Further, the Excess Predicted Proportion (EPP) – a screening performance measure analogous to Highway Safety Manual (HSM), Excess Predicted Average Crash Frequency is proposed for hot zone identification. Using EPP, a statewide screening exercise by the various vehicle types considered in our analysis was undertaken. The screening results revealed that the spatial pattern of hot zones is substantially different across the various vehicle types considered.

**Key words:** multinomial logit fractional split model; traffic crash analysis; macroscopic crash analysis; traffic analysis zones; vehicle type; screening

**1. INTRODUCTION**

The Federal, State, and Local government officials and transportation engineers have been working with consistent efforts in reducing both road crash related fatalities and fatality rates. In the United States, traffic collisions have steadily declined from 2003 to 2011. However, traffic fatalities and fatality rates rose in 2012 with slight drop in 2013 highlighting the challenges faced by the safety community. Overall 29,989 people were reported to be killed in the United States from traffic crashes in 2014 ([NHTSA, 2016](#_ENREF_32)). These facts highlight that there is a need for continued efforts to identify remedial measures to reduce crash occurrences and crash consequences. Towards this end, the traffic safety literature has evolved along two main dimensions: collision frequency analysis and collision severity analysis. The former group of studies is focused on identifying factors that result in traffic collisions while the latter group is concentrated on ameliorating the consequences of traffic crashes (conditional on their occurrences). The current study contributes to traffic safety literature along the first dimension – identify factors that inform us about traffic collision occurrences.

Collision frequency analysis is traditionally undertaken at the microscopic and macroscopic levels. The microscopic safety analysis focuses on roadway entities such as segments, intersections, and corridors (Lee et al., 2017). The studies broadly aim to identify contributing factors for traffic crashes from roadway geometric design, traffic characteristics, and provide specific engineering countermeasures to alleviate traffic collisions. On the other hand, the macroscopic safety analysis relates traffic crashes aggregated at a spatial level (traffic analysis zone (TAZ), census tract or county) with demographic, socioeconomic, built environment, traffic attributes and/or roadway characteristics at a study unit level. While microscopic level analysis is more focused on the engineering design and evaluation, the macroscopic analysis provides a broad spectrum for long-term policy based countermeasures such as enactments of traffic laws, police enforcement, education, and area-wide engineering solutions ([Lee et al., 2014b](#_ENREF_21)). There has been growing recognition within the planning community to incorporate macroscopic models as part of long range transportation plans. For example, the Moving Ahead for Progress in the 21st Century Act (MAP-21) and Fixing America’s Surface Transportation Act (FAST) have emphasized the role of macro-level safety analysis in planning.

While total traffic fatalities and fatality rates show a downward trend, NHTSA statistics indicate that the proportions of motorcycles, bicycles, and pedestrians in fatal crashes have considerably increased whereas the proportion of passenger cars has decreased from 2005 to 2014 ([NHTSA, 2016](#_ENREF_32)). The proportions of motorcyclists and non-motorists involved in fatal crashes has risen from 11% to 14% and 13% to 18%, respectively, since the last decade. Despite these increases, earlier studies in the safety area have predominantly focused either on total crashes or crashes involving passenger cars/trucks or crashes involving non-motorists. These studies provide important information in improving safety situation for different road user groups separately. However, it is also important to examine critical factors contributing to crash occurrences including all road user groups in a single framework, which would allow stakeholders to devise a more general safety conscious planning. Towards that end, the main objective of this study is to explore the proportions of traffic crashes at TAZ level across different vehicle types including both motorized and non-motorized group of road users. Specifically, the current study considers the proportion of crashes by vehicle type as the dependent variable and estimates a TAZ level mixed multinomial fractional split model. The proposed approach will assist transportation planners and engineers in devising safety conscious plans. Specifically, traffic engineers and planners can understand the factors affecting the proportions of crashes by vehicle types from the modeling results, and use these findings to design long-term transportation plans. Furthermore, the predicted proportions of crashes from the proposed model can be used to identify hotspots for each vehicle type. Thus, traffic engineers and planners can proactively provide effective safety countermeasures for the zones with excessive high proportions for specific vehicle types. The vehicle types considered in our analysis include: passenger car[[1]](#footnote-1), van, light truck, medium and heavy truck, bus, motorcycle, bicycle, and pedestrian. The reader would note that the model employed is not similar to the traditional multinomial logit model because the dependent variable in our case is proportion by vehicle type whereas it is a single chosen alternative in the multinomial logit model.

***1.1. Literature Review***

The macro-level safety studies have been conducted by total crashes, crashes by severity levels (such as no injury, minor injury, severe injury, and fatal injury) and crashes by vehicle type (such as motor vehicle, pedestrian, and bicycle). It is beyond the scope of this paper to exhaustively review all the studies in frequency modeling (see [Lord and Mannering (2010)](#_ENREF_24) and Yasmin and Eluru (2016) for a detailed review). In our study, we group literature in the context of our research effort along two main groups: 1) independent frequency models for a single dependent variable or multiple dependent variables; and 2) multivariate count models for the multiple dependent variables are estimated.

In the first group of studies, usually either total number of crashes in the study unit or crashes by vehicle type or severity level are investigated. In some studies, multiple dependent variables are considered in the analysis while ignoring the relationship across the dependent variables. For example, [Noland and Quddus (2004)](#_ENREF_33) developed fixed-effect negative binomial models for bicycle and pedestrian crashes for severe and minor injury. Nevertheless, these models did not account for the possible correlation among dependent variables. [Lee et al. (2013)](#_ENREF_20) estimated a series of negative binomial models for total, severe, driving under the influence (DUI), pedestrian and bicycle crashes based on zero-inflated Poisson (ZIP) framework while ignoring the relationships across the dependent variables. [Abdel-Aty et al. (2013)](#_ENREF_2) analyzed the contributing factors for total, and severe pedestrian crashes, using negative binomial models without considering common unobserved factors across different crash types.

In the second group of studies, multivariate models that recognize the dependencies between multiple dependent variables are estimated. For example, [Song et al. (2006)](#_ENREF_39) analyzed intersection, intersection-related, driveway access, and non-intersection crashes at the county level. The authors developed Bayesian multivariate conditional autoregressive models that can account for the dependencies between the variables and spatial effects. ([Narayanamoorthy et al., 2013](#_ENREF_29)) explored pedestrian and bicycle crash frequencies by injury severity levels. The authors adopted a multivariate model to accommodate jointness in the dependent variables, while considering spatial dependence effects. Furthermore, [Lee et al. (2015b)](#_ENREF_22) estimate multivariate Poisson lognormal models to analyze motor-vehicle, pedestrian, and bicycle crashes. The authors found that the multivariate model accounting for unobserved common factors across the dependent variables outperforms the univariate model that does not take the dependencies between dependent variables into consideration. Recently, [Nashad et al. (2016)](#_ENREF_30) adopted a bivariate copula based modeling approach to examine pedestrian and bicyclist crashes simultaneously.

Based on earlier literature, the factors that are likely to affect motor vehicle crashes are socio-demographic factors, land-use characteristics, roadway-related variables, and traffic characteristics ([Kim et al., 2006](#_ENREF_15); [Lee et al., 2015b](#_ENREF_22)). Based on 0.1 mi2 grid structure, Kim et al. (2006) estimated various regression models based on the grid. The authors identified that population, total job, total economic output, and commercial area had positive relationship with motor vehicle crashes. [Lee et al. (2015b)](#_ENREF_22) analyzed motor vehicle crashes based on TAZs. The authors found that population, total employment, total economic output, commercial area, roadway types, proportion of households without a vehicle, number of accommodation facilities per square mile, and number of traffic signals per mile have a positive effect on motor vehicle crashes. Very few researchers have explored truck-involved crashes at a macro-level. [Pasupuleti and Pulugurtha (2013)](#_ENREF_35) identified several zonal characteristics related to truck crashes. The authors revealed that truck crashes are positively correlated to industrial areas and areas with large residential lots but negatively correlated with highly populated areas.

Several factors that tend to increase bicycle crashes have been identified by researchers. [Noland and Quddus (2004)](#_ENREF_33) found that national health service staff per population, percentage of motorway, percentage of trunk road density, percentage of older vehicle, percentage of households without cars, per capita expenditure on alcohol, population, percentage of population aged 65 or over have a positive relationship with severe bicycle crashes. On the other hand, length of inpatient stay in the hospital, income level, percentage of population aged 45-64 have a negative relationship with severe bicycle crashes. [Kim et al. (2006)](#_ENREF_15) attempted several candidate variables including population, total job, total economic output, hospital, park, commercial area, and school for bicycle crashes but only population was found significant and positively related to bicycle crashes. [Lee et al. (2013)](#_ENREF_20) investigated the residence of bicyclists who were involved in traffic crashes. It was revealed that median age, average travel time to work, household income, and workers in the primary industry field were negatively associated with the number of crash involved bicyclists. On the other hand, Hispanic people, workers commuting by bicycle, urban area, and older buildings were positively associated with the number of bicyclists who were involved in traffic crashes. [Lee et al. (2015b)](#_ENREF_22) developed a multivariate model for motor vehicle-to-vehicle, bicycle-to-vehicle, and pedestrian-to-vehicle crashes. The authors uncovered that vehicle-miles-traveled, population, commuters using bicycle, hotel/motel/timeshare rooms per square mile, employments and school enrollments per square mile, number of traffic signals per mile are likely to increase bicycle crashes whereas proportion of roadway with speed limit of 20 mph or less tends to decrease bicycle crashes. The above mentioned literature suggest that the factors that tend to increase bicycle crashes are population, distance to urban location, employment, school enrollment density, roadway types, number of traffic signals per mile, proportion of households without a vehicle, and household income.

Lastly, some factors that increase the propensity of pedestrian crashes have been found. [Noland and Quddus (2004)](#_ENREF_33) found that total population and percentage of population aged 65 or over are positively correlated whereas percentage of other roads, income, percentage of population aged 45-64 are negatively correlated with severe pedestrian crashes. [Loukaitou-Sideris et al. (2007)](#_ENREF_25) discovered that population, employment density, high traffic volume, commercial/retail land-use, and multifamily residential land-use have a positive effect on pedestrian crashes. In the study of [Wier et al. (2009)](#_ENREF_43), traffic volume, arterials without transit, proportion of mixed land-use (commercial and residential), proportion of commercial land-use, employee and resident populations, and proportion of people below poverty have a positive relationship while land areas, and proportion of population aged 65 or over have a negative relationship with pedestrian crashes. [Lee et al. (2015a)](#_ENREF_19) used the product of population and VMT as an exposure variable in their model. The authors found that proportion of households below poverty level, rail and bus station density, accommodation density, marina/ferry terminal density, and K-12 school density tend to increase pedestrian crashes; but proportion of high-speed roads have a tendency to decrease pedestrian crashes. To the best authors’ knowledge, none of studies have explored motorcycle or bus related crashes at a macro-level. In the data preparation process, we consider the independent variables that have been found significant in the previous studies.

***1.2. Current Study***

In earlier research, the impact of exogenous variables is quantified through the propensity component of count models. The main interaction across different count variables is either ignored (first group) or sought through unobserved effects (studies from second group) i.e. there is no interaction of observed effects across the multiple count models. While this might not be a limitation per se, it might be beneficial to evaluate the impact of exogenous variables in framework that directly relates a single exogenous variable to all count variables of interest simultaneously i.e. a framework where the observed propensities of crashes by vehicle type interact directly. In the traditional count modeling approaches this is not feasible. In this study, an alternative approach to macro-level crash modeling is proposed. Specifically, as opposed to modeling the number of crashes, we adopt a fractional split modeling approach to study the fraction of crashes by each vehicle type for a zone. So for example, in a three-count variable case, the traditional approach would be to adopt a trivariate count model framework with three count equations. In the proposed approach, we adopt a multinomial fractional split model that examines the proportion of crashes (not frequency) by count type with three equations representing the three crash types in a single probabilistic model system. So for a zone, the dependent variable could take the following form – crash type 1: 0.30, crash type 2: 0.25 and crash type 3: 0.45. The fractional split model is used to analyze the proportion by vehicle type across zones as a function of exogenous variables. To be sure, the approach is not a replacement for traditional count based approaches. We believe that the fractional split modeling approach would serve complementary to existing traditional approaches in providing more insights on the impact of exogenous variables on crash proportions.

The fractional split model also provides another advantage. From the review of earlier literature, it is evident that vehicle crashes are usually grouped under one category. However, it is possible that the occurrence of crashes might vary across zones by vehicle type (such as passenger cars, light trucks, medium and heavy trucks, and buses). It is worthwhile to investigate the factors that influence crash occurrence by each vehicle type. However, the addition of different vehicle types would add additional computational burden for count modeling and hence such fine resolution of vehicle types is rarely considered. Within a fractional split model, the additional computation burden associated with adding the fine vehicle type resolution is minimal and thus facilitates considering detailed vehicle type resolution.

The proposed methodology is based on earlier work in econometrics undertaken by [Papke and Wooldridge (1993)](#_ENREF_34). The authors proposed a quasi-likelihood estimation method for binary probit model with a fractional dependent variable. The authors explored 401(K) plan participation rates in two portfolios using their proposed method. The approach was extended to multinomial fractional model by [Sivakumar and Bhat (2002)](#_ENREF_38). The authors analyzed statewide interregional commodity-flow volumes in Texas using the proposed model. [Eluru et al. (2013)](#_ENREF_10) extended the binary probit model by [Papke and Wooldridge (1993)](#_ENREF_34) and proposed a panel mixed ordered probit fractional split model to analyze vehicle operating speed on urban roads in Montreal. Yasmin et al. (2016) employed an ordered version of the fractional split model to investigate crash proportion by injury severity. To be sure, there has been earlier work in safety literature exploring the multinomial fractional split model at the microscopic level. [Milton et al. (2008)](#_ENREF_27) developed a mixed multinomial fractional split model to study injury-severity distribution of crashes on highway segments by using highway-injury data from Washington State. Also several other researchers adopted a fractional split model in the respective fields ([Nam, 2012](#_ENREF_28); [Witter et al., 2012](#_ENREF_44); [Wang & Wolman, 2014](#_ENREF_42)).

Furthermore, we consider the influence of observed spatial effects in our analysis. In the macro-level studies, traffic crashes occurring in a geographic unit are aggregated. The aggregation process might create errors in identifying exogenous variables for the geographic unit. For instance, a crash occurring near or on the boundary of the geographic unit might be strongly related to the neighboring zone than the actual zone where the crash happened, which is the result of arbitrarily defining boundaries. In order to alleviate such geographic unit induced bias, the following two methodologies have been utilized to account for spatial autocorrelations: 1) spatial error correlation effects (unobserved exogenous variables at one location affect dependent variable at the targeted and adjacent zones); and 2) spatial spillover effects (observed exogenous variables at one location impact the dependent variable at both the targeted and adjacent zones) ([Narayanamoorthy et al., 2013](#_ENREF_29)). Some research efforts have accommodated for spatial random error in safety literature ([Huang et al., 2010](#_ENREF_13); [Dong et al., 2014](#_ENREF_7); [Lee, 2014](#_ENREF_17); [Dong et al., 2015](#_ENREF_8); [Lee et al., 2015a](#_ENREF_19); [Lee et al., 2015b](#_ENREF_22); [Dong et al., 2016](#_ENREF_6); [Huang et al., 2016](#_ENREF_14); [Xu et al., 2017](#_ENREF_46)). Nevertheless, using such spatially lagged dependent variable models, specifically for prediction, is of limited use because observed crash at adjacent geographic unit is required as an independent variable in the model. Thus, we adopted a method considering exogenous variables from adjacent zones for accounting for spatial dependency, which was recently suggested ([Cai et al., 2016](#_ENREF_4)). To summarize, in this research we employ a zonal level mixed multinomial fractional split model to investigate the impact of exogenous factors with spatial spillover effects on the proportion of vehicle types in traffic crashes for the state of Florida.

The rest of the paper is organized as follows: Section 2 provides a description of the mixed multinomial fractional split model. Section 3 describes the data collection and sample preparation steps. Section 4 discusses the modeling results and the elasticity effects. A new performance measure for hot zone identification is proposed and screening results are provided in Section 5. Lastly, Section 6 summarizes and concludes the paper.

**2. STATISTICAL FRAMEWORK**

The dependent variable in this study is defined as the proportion of vehicle type in traffic crashes by TAZ. The sum of the proportions across a TAZ is equal to unity and each proportion of vehicle types in traffic crashes ranges between zero and one. Let *ymn* be the fraction of crashes by vehicle type *m* (*m*= 1,2, … , *M; M=8*) in TAZ *n* (n=1, 2, …, *N*). In this paper, the eight vehicle types correspond to passenger car, van, light truck, medium and heavy truck, bus, motorcycle, bicycle, and pedestrian.

(1)

Let the fraction *ymn* be a function of a vector *xmn* of relevant explanatory variables associated with attributes of TAZ *n*.

(2)

where *Gm*(∙)is a predetermined function. The properties specified in Equation (2) for *Gm*(∙) warrant that the predicted fractional crash vehicle types will range between 0 and 1, and will add up to 1 for each TAZ. In this study, a mixed multinomial logit functional form for *Gm* in the fractional split model of Equation 2. Then Equation 2 is rewritten as:

, *m* =1, … , *M* (3)

Given the probability expression above, the quasi likelihood function is written as follows:

(4)

The quasi log-likelihood function for the sample is defined as:

The model estimation is undertaken by maximizing the quasi log-likelihood function based on a routine in Gauss matrix programming language. The readers would note that the coefficient vector includes both mean parameters and standard deviation parameters following a normal distribution.

**3. DATA PREPARATION**

The number of road users involved in crashes for eight vehicle types (passenger car, van, light truck[[2]](#footnote-2), medium and heavy truck, bus, motorcycle, bicycle and pedestrian) were acquired from the Florida Department of Transportation (FDOT) Crash Analysis Reporting System (CARS) for the year 2010 through 2012. The collected crash data by vehicle types were further aggregated based on statewide TAZs (N=8,518) and corresponding crash proportion were computed. Among 8,518 TAZs, there were no crashes reported for 389 TAZs for the study years and hence were excluded from the analysis of our study. The numbers of units by vehicle type in total crashes were aggregated based on TAZ-level. Then the TAZ-based unit counts by vehicle type were converted to the proportions. Table 1 summarizes the descriptive statistics of traffic crash related variables. From the table, we can observe that the proportion of passenger cars (57.5%) is the highest while the proportion of buses (1.1%) is the smallest among all vehicle types considered. The data for the dependent variable is further augmented by socio-demographic, traffic, roadway and commuter travel data from multiple sources such as FDOT CARS/Systems Planning Office/Roadway Characteristics Inventory, Florida Department of Revenue, and U.S. Census Bureau. The prepared independent variables were considered based on previous studies as discussed in the literature review section.

Table : Descriptive statistics of traffic crash related variables

|  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- |
| **Variable (N=8,129)** | **Mean** | **Stdev** | **Min** | **Max** | **% of zero proportion** |
| Total number of crash-involved units based on a TAZ | 146.433 | 214.156 | 1 | 2773 | - |
| Proportion of passenger cars | 0.575 | 0.157 | 0.000 | 1.000 | 1.8% |
| Proportion of vans | 0.061 | 0.056 | 0.000 | 1.000 | 18.3% |
| Proportion of light trucks | 0.264 | 0.139 | 0.000 | 1.000 | 4.2% |
| Proportion of medium & heavy trucks | 0.032 | 0.065 | 0.000 | 1.000 | 33.8% |
| Proportion of buses | 0.011 | 0.030 | 0.000 | 1.000 | 53.6% |
| Proportion of motorcycles | 0.029 | 0.062 | 0.000 | 1.000 | 33.0% |
| Proportion of bicycles | 0.013 | 0.031 | 0.000 | 1.000 | 49.4% |
| Proportion of pedestrians | 0.015 | 0.030 | 0.000 | 0.667 | 45.6% |

Socio-demographic data such as population, family vehicle ownership, hotel/motel/timeshare rooms, employment, and school enrollment were obtained from the Systems Planning Office of the FDOT. Four employment related variables were processed: total employment density, and proportion by employment type - industrial, commercial, and service. Roadway/traffic data were compiled from the FDOT Roadway Characteristics Inventory (RCI) and they were processed at a TAZ level using geographical information systems (GIS). The roadway data include proportion of roadway length by functional classifications (i.e., arterial, collector, and local road), signals per mile, VMT density (VMT normalized by TAZ area), proportion of heavy vehicles, bike lane length density (per square mile), and sidewalk length density (per square mile). Two types of urban classification data were processed. Urban area polygon data (2010) was collected from the U.S. Census Bureau. The urban area data were processed into: 1) proportion of urban areas; and 2) distance to the nearest urban area. Commuting travel data were acquired from the U.S. Census Bureau and the proportions of commuters using specific travel modes (e.g., public transportation, bicycle, walking) were computed for each TAZ. Table 2 exhibits the descriptive statistics of the candidate explanatory variables.

Table : Descriptive statistics of candidate explanatory variables

|  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- |
| **Variable (N=8,129)** | | **Mean** | **Stdev** | **Min** | **Max** |
| Socio-demographic  characteristics | Population density (per sqmi) | 2574.39 | 4032.83 | 0 | 63069 |
| Proportion of young people (15-24 years) | 0.132 | 0.083 | 0 | 1.000 |
| Proportion of elderly people (65 years and older) | 0.171 | 0.116 | 0 | 0.938 |
| Hotel, motel, and timeshare room density (per sqmi) | 173.89 | 951.75 | 0 | 32610 |
| Total employment density (per sqmi) | 1190.3 | 1745.9 | 0 | 31932 |
| Proportion of industrial employments | 0.176 | 0.232 | 0 | 1.000 |
| Proportion of commercial employments | 0.299 | 0.235 | 0 | 1.000 |
| Proportion of service employments | 0.492 | 0.259 | 0 | 1.000 |
| Proportion of families with no available vehicle | 0.095 | 0.123 | 0 | 1.000 |
| School enrollment density (per sqmi) | 803.4 | 6115.3 | 0 | 255147 |
| Median household income (in USD) | 57389 | 24714 | 0 | 215192 |
| Roadway /traffic  characteristics | Arterial proportion of public road mileage | 0.223 | 0.272 | 0 | 1.000 |
| Collector proportion of public road mileage | 0.188 | 0.239 | 0 | 1.000 |
| Local proportion of public road mileage | 0.572 | 0.324 | 0 | 1.000 |
| Number of signals per mile | 3.032 | 88.134 | 0 | 6347.7 |
| Vehicle-miles-traveled density (per sqmi) | 54304 | 187263 | 0 | 11469720 |
| Proportion of trucks | 0.068 | 0.051 | 0 | 0.428 |
| Bike lane length density (per sqmi) | 0.533 | 3.207 | 0 | 153.183 |
| Sidewalk length density (per sqmi) | 3.542 | 10.245 | 0 | 183.373 |
| Presence of transit system (dummy variable) | 0.597 | 0.491 | 0 | 1 |
| Lengths of transit system (mi) | 1.474 | 2.857 | 0 | 59.580 |
| The existence of SIS (Strategic Intermodal System) network | 0.310 | 0.463 | 0 | 1 |
| Number of bus stations | 0.005 | 0.071 | 0 | 1 |
| Number of cargo centers | 0.010 | 0.435 | 0 | 36 |
| Number of commuter rail stations | 0.009 | 0.107 | 0 | 3 |
| Number of train stations | 0.002 | 0.048 | 0 | 1 |
| Land-use  attributes | Proportion of urban areas | 0.732 | 0.424 | 0 | 1.000 |
| Distance to the nearest urban area (mi) | 1.991 | 5.209 | 0 | 44.101 |
| Mixed land-use area (sqmi) | 0.004 | 0.017 | 0 | 0.636 |
| Residential area (sqmi) | 1.265 | 6.020 | 0 | 280.114 |
| Commercial area (sqmi) | 0.100 | 0.225 | 0 | 7.052 |
| Industrial area (sqmi) | 0.045 | 0.428 | 0 | 23.902 |
| Agriculture area (sqmi) | 3.428 | 13.707 | 0 | 227.568 |
| Institutional area (sqmi) | 0.033 | 0.260 | 0 | 11.514 |
| Governmental area (sqmi) | 1.756 | 15.665 | 0 | 748.162 |
| Miscellaneous area (sqmi) | 0.279 | 4.098 | 0 | 338.046 |
| Vacant area (sqmi) | 1.495 | 13.261 | 0 | 702.884 |
| Number of nightclubs, cocktail lounges and bars | 0.234 | 0.757 | 0 | 18 |
| Commuting characteristics | Proportion of commuters using public transportation | 0.024 | 0.043 | 0 | 0.549 |
| Proportion of commuters using bicycle | 0.009 | 0.023 | 0 | 0.309 |
| Proportion of commuters who walk | 0.025 | 0.041 | 0 | 0.449 |
| Indicator variables | District dummy variables (1 to 7) | Percentages | | | |
| District 1 | 11.8% | | | |
| District 2 | 17.1% | | | |
| District 3 | 13.6% | | | |
| District 4 | 12.1% | | | |
| District 5 | 19.3% | | | |
| District 6 | 10.5% | | | |
| District 7 | 15.6% | | | |

**4. RESULTS**

***4.1. Modeling Results***

Table 3 presents the modeling results of the mixed multinomial logit fractional split model. The model estimation results revealed the absence of any significant unobserved parameters. Hence, the model collapsed to a multinomial logit fractional split model. Within the fractional split model, each alternative has a propensity equation. However, for the sake of identification, one of the alternative propensities has to be selected to serve as the base. In our analysis, the model estimation is undertaken considering that the passenger car option serves as the base alternative. Hence, there are no coefficients specific to passenger car in Table 3.

Overall, the model results are intuitive and offer credence to our hypothesis that a fractional split approach to crash proportion modeling is of value. A brief examination of the results highlights substantial differences in the number of significant parameters across alternatives. The propensity for light truck and medium and heavy truck alternatives has more number of significant parameters while pedestrian and bicyclist alternatives have the fewest significant parameters in current study context. Given that the multinomial logit fractional split approach is compensatory in nature (difference of propensities affects proportion), alternatives with smaller proportions (as is the case with pedestrian and bicyclist crashes) are likely to have fewer parameters. In terms of data fit, the additional exogenous variables substantially improve the quasi log-likelihood at convergence (-9738.217) compared to the quasi log-likelihood at zero (-16903.780) and quasi log-likelihood at constants (-9930.663). A log-likelihood ratio test with respect to model at constants yields a test statistic of 384.886 (=2\*(-9738.217 – (-9930.66))). The test statistic is larger than the corresponding chi-square distribution value for 25 additional parameters at any level of significance. The results from the model are discussed subsequently by variable groups.

The reader would note that parameter values in Table 3 correspond to the impact of exogenous variables on the alternative relative to passenger car. Hence, a positive (negative) sign indicates increase (decrease) in the proportion of the alternative relative to the proportion of passenger cars. An insignificant effect implies the variable does not have any differential impact on the alternative relative to the passenger car alternative. The authors also discussed the explanatory variables used in the previous macro-level safety studies. At a first glance, some may think that the findings from prior studies are contradictory to those from this study. However, all the previous studies developed crash count models (by crash types or vehicle types) whereas this study estimated the proportions of vehicle types in traffic crashes, and thus they are not actually conflicting.

Constants

The constants in the model are negative for all alternatives with respect to the base alternative passenger cars. This is expected as the passenger car proportion is substantially larger than other alternatives.

Socio-demographic characteristics

Several sociodemographic variables influence the zonal level proportion of crashes by vehicle types. An increase in the variable, logarithm of population density indicates is associated with a reduction in the proportion of light truck and medium and heavy trucks involved crashes with the magnitude being higher for medium and heavy trucks. In high population density zones, the exposure to light trucks is likely to be much lower (relative to zones with lower population density). The lower exposure might result in lower proportion of light truck involved crashes. Similarly, in high population density zones, the exposure to medium and heavy trucks is lower (relative to other zones) and more importantly the average speed of medium and heavy trucks is also likely much lower thus possibly reducing the proportion of their crash involvement. No prior studies have utilized the population density for the proportion of trucks in traffic crashes at the macro-level. Instead, some studies used the population density as an exploratory variable for total crashes ([Lovegrove & Sayed, 2007](#_ENREF_26); [Lee et al., 2014b](#_ENREF_21); [Xu et al., 2014](#_ENREF_45)), property damage only (PDO) crashes ([Ladron de Guevara et al., 2004](#_ENREF_16)), injury crashes ([Ladron de Guevara et al., 2004](#_ENREF_16); [Lee et al., 2014b](#_ENREF_21); [Xu et al., 2014](#_ENREF_45)), and fatal crashes ([Ladron de Guevara et al., 2004](#_ENREF_16)), bicycle crashes and pedestrian crash count models ([Siddiqui et al., 2012](#_ENREF_37); [Cai et al., 2016](#_ENREF_4)). It was commonly found that the variable had positive effects on those crashes.

The modeling results indicate that a higher proportion of elderly population tends to increase the crash proportion of bicycles and motorcycles (among the crash involved units). Florida has the highest percentage of senior people aged 65 or older and it has experienced consistent in-migration of retired people from other states ([Sperazza et al., 2012](#_ENREF_40)). The large number of retired people might escalate demand for recreational and leisure activities including cycling and motorcycling. Thus, the higher percentage of elderly population may be associated with the larger crash proportion of both bicycles and motorcycles. The elderly population factor has been considered in the several macro-level studies. [Huang et al. (2010)](#_ENREF_13) found that the larger elderly population tends to decrease total and severe crash counts. Also, [Lee et al. (2014a)](#_ENREF_18) accommodated the proportion of elderly population in their study and the variable has a negative relationship with total crash counts.

The variable corresponding to proportion of families with no available vehicle offers interesting results. The variable, on its own, is associated with a reduction in the crash proportion of motorcycles. Motorcycle is mainly a recreational vehicle in Florida and expensive to maintain, hence in such zones the likelihood of ownership as well as the use of motorcycles is likely to be lower. In addition to the TAZ variable, a spatial variable based on the proportion of families with no available vehicle from neighboring TAZs is also considered in our model. The variable is associated with an increase in the crash proportion of buses (among the crash involved units). Families without access to vehicles are captive to public transportation and possibly live in zones with higher access to public transit. Hence, it is not surprising that such zones have higher proportion of bus crashes. On the other hand, some research efforts employed vehicle ownership variable in their studies for other crash types. The higher proportion of households without available vehicle has a propensity to increase bicycle ([Noland & Quddus, 2004](#_ENREF_33)), pedestrian ([Noland & Quddus, 2004](#_ENREF_33); [Lee et al., 2015a](#_ENREF_19)), and total/severe crashes ([Lee et al., 2014b](#_ENREF_21)). [Siddiqui et al. (2012)](#_ENREF_37) used the percentage of households with 0 or 1 vehicle and identified that the variable was positively associated with bicycle and pedestrian crashes. [Quddus (2008)](#_ENREF_36) showed that the logarithm of households without vehicle has a positive impact on fatal and serious injury crashes.

Tourism is an important industry in Florida and it is important to consider tourist presence in traffic safety planning. A surrogate measure for tourist activity is hotel, motel and timeshare facility density. The variable highlights a negative association with the proportion of light trucks in traffic crashes alluding to the possibility that tourist areas are less likely to have larger number of light trucks in the zone. Although no studies have considered the accommodation factor for the proportion of light trucks, some studies found that the hotel units have a positive impact on pedestrian ([Siddiqui et al., 2012](#_ENREF_37); [Lee et al., 2015a](#_ENREF_19); [Lee et al., 2015b](#_ENREF_22); [Cai et al., 2016](#_ENREF_4)), bicycle ([Lee et al., 2015b](#_ENREF_22)), total, severe ([Lee et al., 2014b](#_ENREF_21)), and vehicle-to-vehicle crashes ([Lee et al., 2015b](#_ENREF_22)). On the contrary, [Ng et al. (2002)](#_ENREF_31) showed that the hotel units have a negative effect on fatal crashes.

The employment density variable is associated with a lower likelihood of motorcycle crashes. As motorcycle is mainly used for recreational purposes, it is expected that the use of such vehicles in zones with high employment density is unlikely. In terms of industrial employment, the result indicates that zones with higher industrial employment is likely to involve higher medium and heavy truck crash proportions. The variable is a reflection of increased exposure to medium and heavy trucks in these zones. The findings are further substantiated based on the result of the proportion of trucks estimates. From the previous studies, many researchers have found that employment-related factors have a significant effect on various crash types. An area with larger employment has a tendency to exhibit a greater number of vehicle-to-vehicle and severe crashes ([Hadayeghi et al., 2006](#_ENREF_12)), pedestrian and bicycle crashes ([Siddiqui et al., 2012](#_ENREF_37); [Cai et al., 2016](#_ENREF_4)), and pedestrian crashes only ([Loukaitou-Sideris et al., 2007](#_ENREF_25)). [Levine et al. (1995)](#_ENREF_23) attempted diverse employment variables. The authors found that manufacturing, retail trade, and service employment positively affect crash counts while financial and military employment are negatively related to total crash counts.

Roadway /traffic characteristics

The increased vehicle mileage in the zone is negatively associated with medium and heavy trucks and bus crash proportions. The increased vehicle mileage reflects suburban zones where the exposure to medium and heavy trucks or buses is very low. Hence, the trend is expected. In zones with increased exposure to trucks, a higher incidence of van, light truck and medium and heavy truck crashes is likely with a substantially larger impact on the proportion of medium and heavy trucks. It is also found that the increased vehicle mileages in the neighboring zone have a negative effect on medium and heavy truck proportions. While there have been no macro-level studies exploring the share of trucks in traffic crashes, [Golob and Regan (2003)](#_ENREF_11) uncovered the negative relationship between annual average daily traffic (AADT) level and the probability of truck involvement on urban freeways at the micro-level. This result is quite consistent with the finding from this study.

Land-use attributes

In terms of spatial location of the zone, proportion of urban areas and distance to the nearest urban location exert significant impact on crash proportions. The zones with larger proportion of urban areas, as expected, are likely to have higher incidence of bicycle crashes and lower proportion of light truck and motorcycle crashes. The zones that are farther from urban areas are likely to have a higher proportion of light trucks as these vehicles are more likely to be used in these zones. It was shown that several land-use characteristics have substantial effects on crashes in prior macro-level research studies. [Huang et al. (2010)](#_ENREF_13) showed that the higher level of urbanization is positively related to total and severe crashes, and [Siddiqui et al. (2012)](#_ENREF_37) also showed that an urbanized zone experience the larger number of bicycle crashes.

Moreover, the agriculture area in the neighboring zone has a positive effect on the proportion of crash-involved light trucks. Previous macro-level safety studies have not explicitly shown that the agriculture area per se has a significant impact on traffic safety but some macroscopic safety studies showed that fatal crashes are more likely to occur in rural areas ([Blatt & Furman, 1998](#_ENREF_3); [Stamatiadis & Puccini, 2000](#_ENREF_41); [Clark, 2003](#_ENREF_5)).

Other than urban and land-use variables, we considered FDOT districts as an independent variable in the model. The operations of the FDOT are organized into seven districts and their locations are shown in Figure 1. The result indicates that District 2 has higher crash proportion of medium and heavy trucks while District 4 has a propensity of less light truck crash proportion. District 2 is predominantly characterized by vast rural areas (90.4%) whereas 41.1% area in District 4 is urbanized area.

Commuting characteristics

The zonal commuting characteristics exhibit influence on crash proportions. The zones with high proportion of public transportation commuters are likely to have lower proportion of light truck crashes. The result is an indication of lower light truck ownership in zones with prevalence of public transit ridership. In zones with increased walking commuters, the proportion of pedestrian crashes increases (among the crash involved units). The result is a classic case of higher number of pedestrians in the zone resulting in more pedestrian crashes. In addition, the proportion of commuters using bicycle in adjacent zones has a tendency to increase motorcycle crashes. Several studies have used commuter variables in their studies. [Abdel-Aty et al. (2013)](#_ENREF_2) found that both commuters by public transportation and walking commuters have considerable impacts on total, severe, and pedestrian crashes. [Lee et al. (2014a)](#_ENREF_18) found that zones with higher proportion of commuters using non-motorized modes is more likely to have more total crashes *ceteris paribus*. [Lee et al. (2015b)](#_ENREF_22) exhibited that the logarithm of commuters using bicycle is positively related to the number of bicycle crashes.

***4.2. Elasticity Effects***

The model results from Table 3 provide an indication of how the exogenous variables affect the proportion of crashes involving different vehicle types. However, the exact magnitude of the impact on all alternatives is not easily available. Hence, to evaluate the impact of exogenous variables on all crash proportions, we resort to the computation of elasticity effects. The elasticity effects in our study are computed by evaluating the change in crash proportions in response to increasing the value of significant exogenous variables by 10% (see [Eluru and Bhat (2007)](#_ENREF_9) for more details on elasticity calculations). The results from the elasticity exercise are presented in Table 4. The numbers presented in the table represent the percentage change in the proportion of alternatives in response to a 10% increase in the exogenous variable. For example, the value of elasticity for the variable logarithm of population density of 0.696 represents an increase in passenger car crash proportion with a 10% increase in zonal population density. All the other numbers can be interpreted similarly.

***4.3. Summary***

Based on the results, the following interesting observations may be made. First, the logarithm of population density affects most significantly and negatively the proportion of medium and heavy truck crashes. Second, the proportion of elderly people has a positive effect on bicycle proportions. Third, the proportion of households with no available vehicle affects negatively the proportion of motorcycles whereas the proportion of households with no available vehicle in adjacent zones has the only and positive effect on the crash proportion of buses. Fourth, the logarithm of total employment density has the most significant and negative influence on the motorcycle crash proportion. Fifth, an increase in vehicle-miles-traveled density has a substantial negative impact on the proportion of medium and heavy trucks as well as buses. Sixth, the proportion of trucks has a significant positive association with crash proportion of medium and heavy trucks. Finally, the proportion of commuters who walk has a strong impact on crash proportion of pedestrians.

Table : fractional split multinomial model results for proportion of vehicle types involved in traffic crashes

|  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- |
| **Variables** | | **Van** | **Light truck** | **Medium & heavy truck** | **Bus** | **Motorcycle** | **Bicycle** | **Pedestrian** |
| Constant | | -2.316 (-39.382) | -0.441 (-3.365) | -1.383 (-3.429) | -3.607 (-10.955) | -2.187 (-9.634) | -4.842 (-14.810) | -3.710 (-35.120) |
| Socio-demographic  characteristics | Log of population density (per sqmi) | - | -0.031 (-1.865) # | -0.117 (-3.671) | - | - | - | - |
| Proportion of population aged 65 and more | - | - | - | - | 1.356 (2.625) | 1.906 (2.891) | - |
| Proportion of families with no available vehicle | - | - | - | - | -1.484 (-1.987) | - | - |
| ***Spatial-***Proportion of families with no available vehicle of neighboring TAZs | - | - | - | 2.221 (3.487) | - | - | - |
| Log of hotel, motel, and timeshare rooms density (per sqmi) | - | -0.023 (-1.912) # | - | - | - | - | - |
| Log of total employment density (per sqmi) | - | - | - | - | -0.071 (-1.991) | - | - |
| Proportion of industrial employment | - | - | 0.572 (2.259) | - | - | - | - |
| Roadway /traffic  characteristics | Log of VMT density (per sqmi) | - | - | -0.063 (-2.486) | -0.068 (-2.023) | - | - | - |
| Proportion of heavy vehicles in VMT | 1.074 (2.001) | 1.074 (2.001) | 6.832 (6.718) | - | - | - | - |
| ***Spatial-***Log of VMT density of neighboring TAZs (per sqmi) | - | - | -0.082 (-2.565) | - | - | - | - |
| Land-use  attributes | Proportion of urban areas | - | -0.201 (-1.960) | - | - | -0.795 (-5.043) | 0.848 (2.622) | - |
| Distance to the nearest urban area (mi) | - | 0.010 (1.700) # | - | - | - | - | **-** |
| ***Spatial-***Log of agriculture area of neighboring TAZs (sqmi) | - | 0.028 (1.935) # | - | - | - | - | - |
| District 2 | - | - | 0.306 (2.004) | - | - | - | - |
| District 4 | - | -0.228 (-2.652) | - | - | - | - | - |
| Commuting characteristics | Proportion of commuters using public transportation | - | -2.392 (-3.404) | - | - | - | - | - |
| Proportion of commuters who walk | - | - | - | - | - | - | 2.923 (1.763) # |
| ***Spatial-***Proportion of commuters using bicycle of neighboring TAZs | - | - | - |  | 5.149 (1.743) # | - | - |
| Log-likelihood | Pseudo log-likelihood at 0 | -16903.78 | | | | | | |
| Pseudo log-likelihood at constant | -9930.663 | | | | | | |
| Pseudo log-likelihood at convergence | -9738.217 | | | | | | |

# significant at 90% confidence interval, and all other variables are significant at 95% confidence interval.

Numbers in parentheses are t-values of the estimated coefficients.



**Figure 1: Districts boundaries (1-7) and urban areas (dark colored) in Florida**

Table : Elasticity effects

|  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| **Variables** | | **Passenger car** | **Van** | **Light truck** | **Medium & heavy truck** | **Bus** | **Motorcycle** | **Bicycle** | **Pedestrian** |
| Socio-demographic  characteristics | Log of population density (per sqmi) | 0.696 | 0.698 | -1.201 | -5.448 | 0.694 | 0.688 | 0.696 | 0.695 |
| Proportion of population aged 65 and more | -0.127 | -0.126 | -0.125 | -0.118 | -0.122 | 2.459 | 3.815 | -0.127 |
| Proportion of families with no available vehicle | 0.032 | 0.032 | 0.031 | 0.031 | 0.037 | -1.082 | 0.032 | 0.034 |
| ***Spatial-***Proportion of families with no available vehicle of neighboring TAZs | -0.036 | -0.035 | -0.029 | -0.031 | 3.075 | -0.025 | -0.038 | -0.042 |
| Log of hotel, motel, and timeshare rooms density (per sqmi) | 0.086 | 0.085 | -0.235 | 0.062 | 0.088 | 0.071 | 0.096 | 0.089 |
| Log of total employment density (per sqmi) | 0.111 | 0.111 | 0.115 | 0.113 | 0.106 | -3.780 | 0.111 | 0.110 |
| Proportion of industrial employment | -0.038 | -0.040 | -0.051 | 1.272 | -0.038 | -0.050 | -0.028 | -0.037 |
| Roadway /traffic characteristics | Log of VMT density (per sqmi) | 0.221 | 0.226 | 0.236 | -4.876 | -5.640 | 0.232 | 0.211 | 0.222 |
| Proportion of heavy vehicles in VMT | -0.415 | 0.262 | 0.237 | 5.991 | -0.407 | -0.536 | -0.334 | -0.407 |
| ***Spatial-*** Log of VMT density of neighboring TAZs (per sqmi) | 0.262 | 0.270 | 0.313 | -8.434 | 0.267 | 0.311 | 0.228 | 0.258 |
| Land-use attributes | Proportion of urban areas | 0.031 | 0.031 | -0.062 | 0.031 | 0.028 | -0.361 | 0.405 | 0.030 |
| Distance to the nearest urban area (mi) | -0.058 | -0.062 | 0.178 | -0.147 | -0.061 | -0.109 | -0.028 | -0.056 |
| ***Spatial-***Log of agriculture area of neighboring TAZs (sqmi) | -0.063 | -0.068 | 0.196 | -0.179 | -0.066 | -0.132 | -0.020 | -0.060 |
| District 2 | -0.981 | -1.009 | -1.147 | 31.370 | -1.019 | -1.146 | -0.874 | -0.972 |
| District 4 | 5.625 | 5.659 | -15.828 | 6.344 | 5.539 | 6.112 | 5.332 | 5.574 |
| Commuting characteristics | Proportion of commuters using public transportation | 0.126 | 0.126 | -0.349 | 0.113 | 0.157 | 0.097 | 0.128 | 0.134 |
| Proportion of commuters who walk | -0.016 | -0.016 | -0.013 | -0.012 | -0.021 | -0.013 | -0.017 | 0.953 |
| ***Spatial-***Proportion of commuters using bicycle of neighboring TAZs | -0.013 | -0.013 | -0.011 | -0.010 | -0.014 | 0.425 | -0.014 | -0.015 |

**6. HOT ZONE IDENTIFICATION FOR SPECIFIC VEHICLE TYPE**

To facilitate the application of multinomial fractional split model for hot zone identification, we propose a measure based on the predicted proportions. The measure is analogous to the Highway Safety Manual (HSM) ([AASHTO, 2010](#_ENREF_1)) performance measure for identifying crash hotspots. The HSM approach employs Excess Predicted Average Crash Frequency Using Safety Performance Functions. The measure is calculated by subtracting the predicted crash frequency from the observed crash frequency. When the excess predicted average crash frequency is greater than zero, a zone experiences more traffic crashes than predicted. On the other hand, when the excess predicted average crash frequency is less than zero, a zone experience fewer traffic crashes than predicted.

Similar to this method, we propose the Excess Predicted Proportion (EPP) for a macroscopic screening performance measure, which is the difference between the observed and predicted proportion of each vehicle type for a TAZ.

(5)

where, *EPPmn* is the Excess Predicted Proportion of crash vehicle type *m* at TAZ *n*. *P*(*obs*)*mn* is the observed proportion of crash vehicle type; and *P*(*prd*)*mn* is the predicted proportion of crash vehicle type estimated from the fractional split multinomial logit model (Table 3). When EPP exceeds zero, the type proportion for that TAZ is higher than predicted. In contrast, when EPP is smaller than zero, the type proportion for that TAZ is lower than predicted.

The EPP approach is slightly different from the earlier count based approach. Because we deal with proportions that add up to 1 for a TAZ, a positive EPP for one vehicle type automatically causes a negative EPP for at least another vehicle type. Hence, directly identifying the zones with positive EPP as hot zones will not be appropriate. Because every zone will be a hot zone for at least one mode (unless EPP is exactly 0 for all modes). Hence, after computing the measure across the entire sample for all vehicle types, a vehicle type specific hot zone (H) is identified based on the top 10 percentile ranking of TAZs for that vehicle type. The other zones are labeled as normal (N). An illustration of EPP computation for the passenger car mode is provided in Table 5. The TAZs are arranged in descending order of EPP values. Clearly, TAZs #5762, #6944, and #7882 were classified as a hot zone for passenger cars since their EPPs are within top 10% in the study area. Therefore, it is recommended to focus on passenger car-involved crashes at these zones, since the proportion of crash-involved passenger cars far exceeded the expected proportion of crashes. The EPP measure is a useful tool for policy makers and practitioners to understand potential safety risks by various vehicle types, and provide an adequate safety countermeasure solely for the problematic vehicle types. The zones identified as risky might be very different from zones identified as risky using the HSM hot zone identification process. We believe that this provides an alternative paradigm for safety analysis. The reader would note that analogous measures to excess expected proportion can also be computed using the proposed model.

Table : Examples of screening results of passenger cars

|  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- |
| Rank based on passenger car EPP | TAZ No | Passenger car | | | | |
| *Pobs* | *Ppred* | EPP | percentile | classification |
| 1 | 5762 | 1 | 0.317 | 0.683 | 0.01% | H |
| 2 | 6944 | 1 | 0.373 | 0.627 | 0.02% | H |
| : | : | : | : | : | : | : |
| 812 | 7882 | 0.800 | 0.670 | 0.130 | 9.99% | H |
| 813 | 1474 | 0.617 | 0.487 | 0.130 | 10.00% | N |
| : | : | : | : | : |  | : |
| 8129 | 6241 | 0 | 0.649 | -0.679 | 100.00% | N |

To further illustrate the value of the proposed framework, we identify hot zones across the state of Florida for all vehicle types. To clarify the presentation a figure that depicts the main urban areas in the State of Florida is provided (see Figure 1). In Figure 2 statewide screening results for all 8 vehicle types are provided. An observation of Figure 2 clearly shows that the spatial pattern of hot zones varies considerably across the various vehicle types. It is interesting to note that the spatial pattern of hot zones for passenger cars, vans, light trucks vary substantially across the region. For passenger cars the hot zones correspond to urban areas. The spatial distribution of van specific hot zones does not follow any particular pattern. On the other hand, the spatial distribution of light truck specific hot zones clearly indicates a predisposition for rural areas. For medium and heavy trucks, the hot zones are concentrated in central and south rural areas. The overall bus hot zones tend to be located in urban and suburban areas with minor exceptions. Motorcycle hot zones seem typically located in rural areas. The hot zones for bicyclists are apparently found in large metropolitan areas whereas rural areas are relatively safe for bicyclists. Lastly, the pedestrian hot zones are generally located in urban and suburban areas, where many residential areas are located.

|  |  |  |  |
| --- | --- | --- | --- |
|  |  |  |  |
|  |  |  |  |

Figure : Statewide screening results for each vehicle type

**7. SUMMARY AND CONCLUSION**

In recent years, there is growing recognition that incorporating macroscopic traffic safety analysis into the long-term transportation planning process is beneficial. The macroscopic safety analysis relates traffic crashes aggregated at a spatial level (traffic analysis zone (TAZ), census tract or county) with demographic, socioeconomic, and zonal level traffic/roadway characteristics. The macroscopic analysis can provide a broad spectrum perspective and long-term policy suggestions. The macro-level safety studies have been conducted by total crashes, crashes by severity level (such as no injury, minor injury, severe injury, and fatal injury) and crashes by vehicle type (such as motor vehicle, pedestrian, and bicycle). The range of approaches employed in literature includes univariate count frequency approaches for single dependent variable and multivariate count frequency models for multiple dependent variables. It has been recognized that in the presence of multiple dependent variables, multivariate approaches are appropriate.

In this study, we propose an alternative approach to examining multiple crash dependent variables. Specifically, a mixed multinomial logit fractional split model that examines the proportion of crash by vehicle type is developed. In this model, each alternative proportion is associated with a propensity. The proportion allocated to an alternative is probabilistically determined based on the alternative propensity as well propensity of all other alternatives. Thus, in this approach, exogenous variable effects directly affect all alternatives. In count frequency models, such interactions are absent. Thus, the proposed approach allows an alternate mechanism to examine multiple crash dependent variables. Further, the approach is well suited to accommodate a large number of alternatives without a sizable increase in computational burden. On the contrary, in a count modeling approach developing multivariate models for large number of dependent variables is computationally and methodologically challenging. The advantage is illustrated in our study by considering eight vehicle types in our analysis – passenger car, van, light trucks, medium and heavy trucks, bus, motorcycle, bicycle and pedestrian. The modeling approach allows us to identify and quantify the factors affecting crash proportions at the macro-level.

The proposed mixed multinomial logit fractional split model was estimated using socio-demographic, traffic, land-use, and commuting data at a Traffic Analysis Zone level using data compiled from multiple sources in Florida. The modeling results clearly highlight the applicability of the proposed approach for crashes involving different vehicle type analysis. The modeling results revealed that the impact of explanatory variables varies significantly across different vehicle types. For example, medium and heavy truck crashes are influenced by both socio-demographic (such as population density and industrial employment) and traffic characteristics (such as vehicle-miles-traveled density and proportion of trucks). On the other hand, non-motorized vehicle types (i.e. bicycle and pedestrian) had only one variable related to urban location or commuting characteristics. Subsequently, to quantify the impact of variables across the alternatives, elasticity effects were computed and presented.

To illustrate the applicability of the proposed framework for screening purposes, we also proposed the Excess Predicted Proportion (EPP) measure that computes the difference between the observed and predicted proportion of each vehicle type involved crashes in a zone. Hence, EPP, analogous to the Highway Safety Manual (HSM) Excess Predicted Average Crash Frequency, allows us to identify unsafe or hot zones. Based on this measure, a statewide screening exercise by the various vehicle types considered in our analysis was undertaken. The screening results revealed that the spatial pattern of hot zones is substantially different across the various vehicle type crashes. The screening exercise clearly illustrates the value of the proposed approach.

Overall, the paper demonstrated the application of a mixed multinomial fractional split model for crash proportion modeling as statewide screening purposes (using EPP). The findings from our study are useful for policy makers and practitioners to understand potential safety risks by various transportation vehicle types and provide appropriate and effective policy-based countermeasures by vehicle types. The proposed approach is quite flexible and can be adopted to examine crash proportions by other crash attributes, such as collision type, severity level and temporal classification.

To be sure, the study is not without limitations. The consideration of crash proportions (as opposed to crash counts) might result in incorrect identification of problematic zones under rare circumstances (with very low number of crashes). To address this limitation, future studies can consider coupling a crash count model (for total crashes) with a fractional split model to ensure that proportion and frequency are simultaneously considered. Finally, due to the compensatory nature of the model structure an increase (or decrease) in crash proportion for an alternative is associated with decrease (or increase) in crash proportions from other alternatives. This needs to be recognized while developing and employing fractional split models for crash proportions.

**ACKNOWLEDGMENT**

The authors wish to thank the Florida Department of Transportation for providing data and funding for this study.

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1. Passenger cars include sedans and sport utility vehicles (SUV). [↑](#footnote-ref-1)
2. The reader would note that not only passenger car but also van and light truck represent vehicles used for passenger transport. [↑](#footnote-ref-2)